## Gaussian Markov random fields for MRF sampling

(and inference)

Jean-Baptiste Courbot<sup>1</sup> & Hugo Gangloff<sup>2</sup>

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<sup>1</sup> IRIMAS - UR 7499 – Université de Haute-Alsace, Mulhouse, France

<sup>&</sup>lt;sup>2</sup> Université Paris-Saclay, AgroParisTech, INRAE, UMR MIA Paris-Saclay, Palaiseau, France





I. Introduction: MRF, GMRF

II. Sampling

II.1 Method

II.2 Some results

III. Inference 🛦

IV. Next steps

Introduction: MRF, GMRF

## MRF in image processing

Let S be the lattice of n sites in an image.  $\mathbf{X} = \{X_s\}_{s \in S}$  is a MRF if and only if,  $\forall s \in S$ :

$$p(x_s|\mathbf{x}_{S\setminus s})=p(x_s|\mathbf{x}_{N_s})$$

where  $N_s$  denotes the neighborhood of the s site.

The Hammersley-Clifford theorem<sup>1</sup> allows to write, considering only pairwise site interactions  $\psi$ :

$$p(\mathbf{x}) = rac{1}{\gamma} \exp \left( -\sum_{\mathbf{s} \in \mathcal{S}} \sum_{\mathbf{s}' \in \mathcal{N}_{\mathbf{s}}} \psi(\mathbf{x}_{\mathbf{s}}, \mathbf{x}_{\mathbf{s}'}) 
ight).$$

Computing  $\gamma > 0$  is intractable in general, so realizations  $\mathbf{X} = \mathbf{x}$  are obtained through iterative sampling techniques, such as Gibbs sampling<sup>2</sup>.

<sup>&</sup>lt;sup>1</sup>Clifford and Hammersley, "Markov fields on finite graphs and lattices", 1971.

<sup>&</sup>lt;sup>2</sup>Geman and Geman, "Stochastic relaxation, Gibbs distributions, and the Bayesian restoration of images", 1984.

#### **GRF and GMRF**

 ${f Z}$  is a Gaussian random field on  ${\cal S}$  iff  ${f Z} \sim {\cal N}({m \mu},{f \Sigma})$ , with  ${m \mu} \in \mathbb{R}^n$  and  ${f \Sigma} \in \mathbb{R}^{n imes n}$  a covariance matrix.

Upon known conditions on  $\Sigma^3$ , GRF are also Markovian, and are then referred to as Gaussian Markov random fields (GMRF). Depending on  $\Sigma$ , GMRF can be sampled with:

- Cholesky decomposition: S is small.
- Fourier sampling: S is a torus and  $\Sigma$  is circulant.
- Spectral sampling: Σ belongs to an extended Gneiting class of covariances<sup>4</sup>.

<sup>&</sup>lt;sup>3</sup>Rozanov, Markov random fields, 1982, p. 120.

<sup>&</sup>lt;sup>4</sup> Allard et al., "Simulating space-time random fields with nonseparable Gneiting-type covariance functions", 2020.

## Our proposition

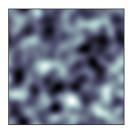
Sampling a GMRF in  $\mathbb{R}^n$  is computationally efficient:

- We propose to use GMRF as proxies for discrete fields in  $\Omega^n$
- What are the properties of the resulting field?
- How can this help for inference?

## Sampling

#### The initial idea

'If you look at a thresholded GMRF, it looks like a MRF"



(a)  $\mathbf{Z} = \mathbf{z}$ , GMRF real.



(b)  $\{z_s \geq 0\}_{s \in \mathcal{S}}$ 



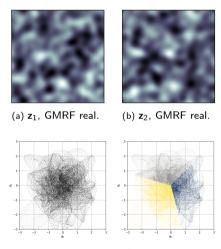
(c) MRF real.

## Moving to K > 2 classes

How to "threshold" to have K > 2 classes, ensuring class balance ?

- we need  $Z_s$  to lie in a higher dimension
- Design **Z** as a (K-1)-valued GMRF, such that now  $\mathbf{z} \in \mathbb{R}^{(K-1)n}$ , and at each site  $s \ \mathbf{z}_s \in \mathbb{R}^{K-1}$
- Example in the K = 3 classes cases

Thus, in dimension K-1 we can split into K balanced classes.



(c) location of each  $\textbf{z}_s$  (d) Splitting in 3 in  $\mathbb{R}^2$  classes

## Formalizing (I)

We need to formalize:

- the splitting within  $\mathbb{R}^{K-1}$
- a proper writing of X given Z

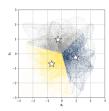


Fig. 1: Unit 2-simplex and splitting

#### Definition (Unit simplex)

A unit P-simplex is a regular simplex belonging in  $\mathbb{R}^P$ , whose P+1 vertices lie on a unit sphere. From<sup>5</sup> we take its vertices  $\mathbf{v} \in \mathbb{R}^P$  as  $\mathbf{v}_{P+1} = \frac{-1}{\sqrt{P}} \mathbf{1}$  and :

$$\mathbf{v}_j = \sqrt{\frac{P+1}{P}}\mathbf{e}_j - \frac{\sqrt{P+1}-1}{P\sqrt{P}}\mathbf{1} \ \forall 1 \leq j \leq P$$

with  $\mathbf{1} \in \mathbb{R}^P$  a vector of ones and  $\mathbf{e}_i \in \mathbb{R}^P$  the j-th basis vector.

<sup>&</sup>lt;sup>5</sup>Anderson and Thron, "Coordinate Permutation-Invariant Unit N-Simplexes in N dimensions", 2021.

## Formalizing (II)

#### Definition (Gaussian Unit-simplex Markov random field (GUM))

Let K be the number of classes to sample from, and  $\mathbf{U}_{K-1}$  the K vertices of a unit (K-1)-simplex. Let also  $\mathbf{Z} \sim \mathcal{N}(\mathbf{0}, \mathbf{\Sigma})$  be a GMRF taking values in  $\mathbb{R}^{n(K-1)}$ , such that  $\mathbf{Z} = \{\mathbf{Z}_s\}_{s \in \mathcal{S}}$  and  $\mathbf{Z}_s$  takes values in  $\mathbb{R}^{K-1}$ .

We define  $\phi_{K,c} \colon \mathbb{R}^{n(K-1)} \mapsto \mathbb{R}^n$  such that:

$$\phi_{K,c}(\mathbf{Z}) = \sum_{i=1}^K \omega_i \pi_i^c(\mathbf{Z})$$

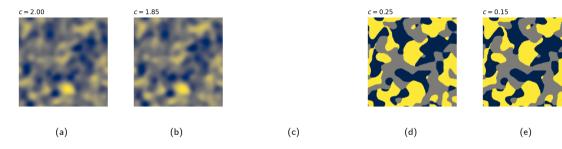
 $\pi_i^c$  indicates, site-wise, the distance between **Z** and the i-th vertices  $\mathbf{v}_i$  of  $\mathbf{U}_{K-1}$ , such that  $\forall s \in \mathcal{S}$ :

$$\pi_i^c(\mathbf{Z}_s) = \frac{\exp(-c^{-2}\|\mathbf{Z}_s - \mathbf{v}_i\|^2)}{\sum_{k=1}^K \exp(-c^{-2}\|\mathbf{Z}_s - \mathbf{v}_k\|^2)}$$

with c > 0,  $\mathbf{v}_k, \mathbf{v}_i \in \mathbf{U}_{K-1}$  unit simplex vertices, and  $\omega_i \in \Omega \subset \mathbb{N}$ .  $\phi_{K,c}(\mathbf{Z})$  is named a *GUM random field*.

#### **GUM**: illustration

A look at the values taken by  $\phi_{K,c}(z)$ :



Property 1: **Z** being a GMRF, then  $\phi_{K,c}(\mathbf{Z})$  is also a Markov field.

Property 2: when  $c \to 0$ ,  $\phi_{K,c}(\mathbf{Z})$  get close to a mixture of Dirac masses:

$$\phi_{K,c}(\mathbf{Z}) \xrightarrow[c \to 0]{} \sum_{i=1}^{K} \omega_i \delta_{\left[\|\mathbf{Z} - \mathbf{v}_i\|_2 \le \|\mathbf{Z} - \mathbf{v}_k\|_2, \ \forall \mathbf{v}_k \in \mathbf{U}_{K-1}\right]}$$

#### Discrete GUM

Property 2: when  $c \to 0$ ,  $\phi_{K,c}(\mathbf{Z})$  get close to a mixture of Dirac masses:

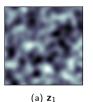
$$\phi_{K,c}(\mathbf{Z}) \underset{c \to 0}{\longrightarrow} \sum_{i=1}^{K} \omega_{i} \delta_{\left[\|\mathbf{Z} - \mathbf{v}_{i}\|_{2} \leq \|\mathbf{Z} - \mathbf{v}_{k}\|_{2}, \ \forall \mathbf{v}_{k} \in \mathbf{U}_{K-1}\right]}$$

Rewording  $\lim_{c\to 0} \phi_{K,c}(\mathbf{Z}) = \mathbf{X} = \{X_s\}_{s\in\mathcal{S}}$ , we have  $\forall s\in\mathcal{S}$ :

$$X_s = \omega_{k^*}$$
 with  $k^*$  chosen such that  $\mathbf{v}_{k^*} = rg \min_{\mathbf{v} \in \mathbf{U}_{K-1}} \|\mathbf{Z}_s - \mathbf{v}\|_2$ 

Some insights 🛦

- X should be Markovian too
- related to a multivariate sigmoid and / or a transformed gaussian random field





(b)  $\mathbf{z}_2$ 





(c) Location & labels

(d) **x** 

#### **Numerical results**

 $\label{lem:lem:python \& JAX, in CPU and GPU. Available on Github: $$ $$ https://github.com/HGangloff/mrfx $$$ 

We want to assess:

- sampling speed
- statistical properties

#### We evaluate:

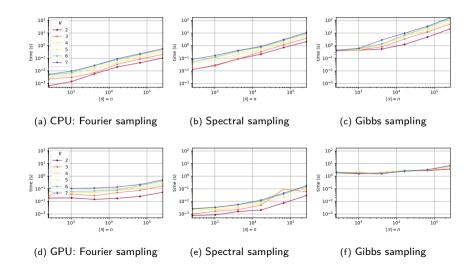
- DGUMS based on Fourier GMRF sampling<sup>6</sup>
- and on spectral GMRF sampling<sup>7</sup>
- MRF with chromatic Gibbs sampling<sup>8</sup>

<sup>&</sup>lt;sup>6</sup>Rue and Held, Gaussian Markov random fields: theory and applications, 2005.

<sup>&</sup>lt;sup>7</sup>Allard et al., "Simulating space-time random fields with nonseparable Gneiting-type covariance functions", 2020.

<sup>&</sup>lt;sup>8</sup>Gonzalez et al., "Parallel Gibbs sampling: From colored fields to thin junction trees", 2011.

## Numerical results: sampling speed

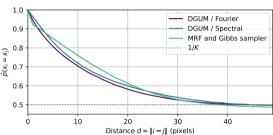


## Numerical results: statistical properties

Pointwise measures, denoting  $\pi_k = p(x_s = k)$ 

	К	2	3	4	
DGUM / Fourier	$\hat{\pi}_0$	0.4924	0.3278	0.2369	, (× = ×)
	$std(\hat{\pi}_0)$	0.0650	0.0497	0.0533	
DGUM / Spectral	$\hat{\pi}_0$	0.5036	0.3296	0.2534	
	$std(\hat{\pi}_0)$	0.0935	0.0780	0.0711	
Gibbs sampling	$\hat{\pi}_0$	0.4980	0.3470	0.2480	
	$std(\hat{\pi}_0)$	0.0399	0.0581	0.0505	

#### Pairwise statistics:



# Inference 🖄

## Inferring with GUMs

Observation model: we assume  $\mathbf{y}$  is obtained as a noisy version of  $\mathbf{x}$ . For instance: class-wise parameters changes within Gaussian or Poisson distributions.

We conjecture that inferring with a GUM prior is not tractable.

The GUM inverse problem relies on:

$$p(\mathbf{y}, \mathbf{x}) \propto p(\mathbf{y}|\mathbf{x})p(\mathbf{x}) = p(\mathbf{y}|\mathbf{x}) \int_{\mathbb{R}^{n(K-1)}} p(\mathbf{x}|\mathbf{z}) d\mathbf{z}$$

- not that simple
- too much information is lost in p(x|z)



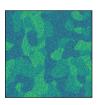
(a) **z**<sub>1</sub>



(b)  $\mathbf{z}_2$ 



(c) x



(d) y

## The problems & solutions: inference with Kernel MLE

 $\forall s$  we can write a likelihood:

$$L_k(y_s) = p(y_s|x_s = \omega_k)$$

We then design a maximum of kernel MLE,  $\forall s$ :

$$\hat{\mathbf{x}}_{
ho} = rg \max_{k} \ ig(\mathbf{\Sigma}_{
ho} L_{k}(\mathbf{y})ig)_{s}$$

with  $\mathbf{\Sigma}_{\rho}$  depending on a covariance function of range parameter  $\rho.$ 

Then: how to determine  $\rho$ ?















## The problems & solutions (II): optimization problem

How to determine  $\rho$  ?  $\rightarrow$  a MAP estimator:

$$\rho^* = \argmax_{\rho > 0} \; p(\mathbf{y}|\hat{\mathbf{x}}_\rho) p(\hat{\mathbf{x}}_\rho)$$

Let us assume a Potts-like potential for  $\mathbf{x}$ , assuming  $\lambda > 0$ :

$$p(\mathbf{x}) \propto \exp\left(-\lambda U(\mathbf{x})\right)$$
, and  $U(\mathbf{x}) = \sum_{s \in \mathcal{S}} \sum_{s' \in N_s} \mathbb{1}_{\{x_s = x_{s'}\}}$ 

Then the MAP estimator can be rewritten:

$$\underset{\rho>0}{\arg\min} \ -\log(p(\mathbf{y}|\hat{\mathbf{x}}_{\rho})) + \lambda U(\hat{\mathbf{x}}_{\rho}) \tag{$\mathcal{P}_{\lambda}$}$$

This can be solved quite easily numerically, for a given  $\lambda$ , because computing  $\hat{\mathbf{x}}_{\rho}$  is fast.

Then: how to determine  $\lambda$ ?

## The problems & solutions (III): homotopy approach $ilde{\mathbb{A}}$

How to determine  $\lambda$  ?  $\rightarrow$  Proximity to sparse optimization problems:

- Solutions of  $(\mathcal{P}_{\lambda})$  lies on the Pareto frontier
- Homotopy methods probe this frontier until a solution is found

Purpose : we can trade the choice for  $\lambda$  with the choice for a  $\ell_2\text{-related}$  parameter.

Fig. 2: Pareto frontier: solutions to 
$$(\mathcal{P}_{\lambda})$$
 are under-optimal above the curve and unattainable below.

$$\left\| \frac{\mathbf{y} - \boldsymbol{\mu}_{\hat{\mathbf{x}}_{\rho}}}{\sigma_{\hat{\mathbf{x}}_{\rho}}} \right\|_{2}$$
 such that  $\rho$  is solution to  $(\mathcal{P}_{\lambda})$   $(\mathcal{P}_{\mathrm{hom}})$ 

Rephrasing : correctness of solution found for  $\lambda$  with respect to the known noise parameter (mean, variance) gathered under  $\Theta$  ?

## Inference: wrapping up 🛕

#### Several layers of nested problems:

- The core problem: an estimator  $\hat{\mathbf{x}}_{\rho}$  that is computationally cheap, for a given  $\rho$
- $\rho$  is found solving  $(\mathcal{P}_{\lambda})$  for a given  $\lambda$
- $\lambda$  is found solving  $(\mathcal{P}_{hom})$  given noise parameters  $\Theta$
- **O** is estimated by an alternating scheme (not detailed here)

#### Preliminary results:

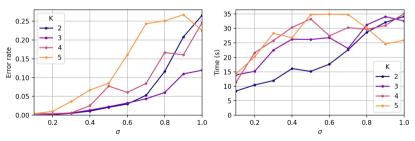


Fig. 3: Error rate (left) and computation time(right) under Gaussian noise with means at  $\{1,2,\ldots,K\}$  and varying standard deviation  $\sigma$ .



### **Next steps**

#### Sampling:

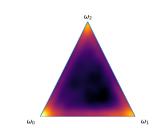
- Properties of X
- Class unbalance

#### Inference:

- Relation between inference formalism and GUMS (if any)
- Formalize the problem nesting & relation to sparse optimization
- Numerical results & real microscopy images

#### Representations:

- The probability simplex
- Uncertainty quantification



(a) Pointwise DGUM density in the probability simplex (K=3).



(b)  $\hat{\mathbf{x}}_{
ho}$ 



(c) Tentative uncertainty

#### Resources

Take-home message: to avoid Gibbs-induced computational costs, we propose:

- a fast sampling method that mimic Potts models
- a fast nested inference method, suitable for unsupervised inverse problems

#### Resources so far:

- Sampling was described in SSP 2025<sup>9</sup>
- Sampling code is available<sup>10</sup>

#### In progress:

- journal paper to wrap this up
- related: a tutorial on Markov models in image processing (with Julien Stoehr)

<sup>&</sup>lt;sup>9</sup>Courbot and Gangloff, "Gaussian Unit-simplex Markov random fields as a fast proxy for MRF sampling", 2025.

<sup>10</sup>https://github.com/HGangloff/mrfx